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ORIGINAL ARTICLE

Who bears labour taxes and social contributions? A meta-analysis approach

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Abstract In this paper we apply the meta-regression technique to survey the empirical literature on the economic incidence of labour taxes and social security contributions. In particular, we focus on the effects of taxation on wages to test the conventional view that employees bear the burden due to lower net wages. Based on 52 empirical papers, we find that economic institutions, the tax wedge definition, and the temporal focus significantly affect the results. In the long run, workers bear between two thirds of the tax burden in Continental and Anglo-Saxon economies, and nearly 90 % in the Nordic economies. However, despite the numerous set of controlling variables, a significant part of the variability of the empirical literature remains unexplained.

Keywords Labour taxes · Incidence · Meta-analysis

JEL Classification C83 · E24 · H22

1 Introduction

The reduction in labour taxes is a widespread policy recommendation for raising employment (see, for instance, the seminal reports from the [European Commission 1994](#); [OECD 1994](#)). In broad terms, labour taxes (i.e. personal income tax and social security contributions) drive a wedge between labour costs and net wages and have a negative effect on labour supply, structural employment and hours worked. From the academia, [Prescott \(2004\)](#) triggered the debate by attributing all the difference in

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labour utilization between the US and Europe to direct taxes. This author calibrated a labour supply model and found out that the divergence in hours worked per week among the working-age population since the 1970s between the US and France and Germany could be explained by differences in marginal tax rates on labour. In a similar line of research, [Coenen et al. \(2008\)](#) employed a calibrated version of the New Area-Wide Model developed at the ECB to simulate the effects of establishing a ‘US fiscal system’ in the euro area. The analysis showed that the reduction in employer social contributions to the levels prevailing in the US (from 21.9 to 7.1 % of labour costs) may increase output by more than five points, hours worked by more than 6 % and real wages slightly less than 13 %. These results generated a lively debate.¹

Based on this assumption, since the 1990s many European governments have followed this tax reform path, cutting social security payroll taxes for cyclical and structural reasons. For instance, since 1997 Spain has cut social security payroll taxes for permanent contracts and for population groups affected by long-term unemployment. Recently, the new government has announced a reduction in employer social contributions (1 p.p. in 2013 and another point in 2014), compensated by an increase in value added taxation. Since 2000, France encouraged the transition to the 35-hour week with lower employer social security contributions. In 2007 and 2008 Germany introduced cuts in unemployment insurance contributions, financed by a higher value added tax rate. Finally, in the midst of the recent international crisis, the US [Congressional Budget Office \(2008\)](#), for instance, considered the reduction in social contributions to be one of the most effective measures for responding to short-term economic weakness, albeit subject to lags and uncertainty.

Despite these common developments, there remains a significant fiscal gap within OECD countries. As [Fig. 1](#) shows, the direct tax wedge (income tax, employee and employer payroll taxes) for a two-earner household with two children is well over 40 % in France, Italy and Germany, while Anglo-Saxon economies and Japan limit the burden to approximately 25–30 %.

The economic effects of taxation depend ultimately on the long and short-term economic incidence, i.e. on who really bears the burden. In the case of employer social contributions, they can be borne by the employers (ultimately reducing the firm’s profits), they can be shifted backwards to employees (reducing net wages), or they can be shifted forward to consumers (increasing the price level). Most of the previous papers calibrate this effect.

Empirical literature shows mixed results. In a classic survey, [Hamermesh \(1993\)](#) analysed 15 seminal studies on the economic incidence of payroll taxes, mainly social security contributions. The author rejected any robust conclusion, not even a consensus interval: results ranged from full to null shifting. By surveying recent studies, [Arpaia and Mourre \(2005\)](#) confirmed that taxation increased unemployment but they also highlighted the complexity of its interactions with other labour market and economic

¹ These findings are supported by [Ohanian et al. \(2006\)](#) for a wider sample of economies. Other authors have suggested complementary or alternative explanations. [Nickell \(2003\)](#) points to social protection rules, [Blanchard \(2004\)](#) to preferences, [Alesina et al. \(2006\)](#) to the role of labour protection and unions, [Rogerson \(2007a\)](#) to taxes and technological progress and [Ljungqvist and Sargent \(2007\)](#) and [Rogerson \(2007b\)](#) to social benefits and the use of revenues.

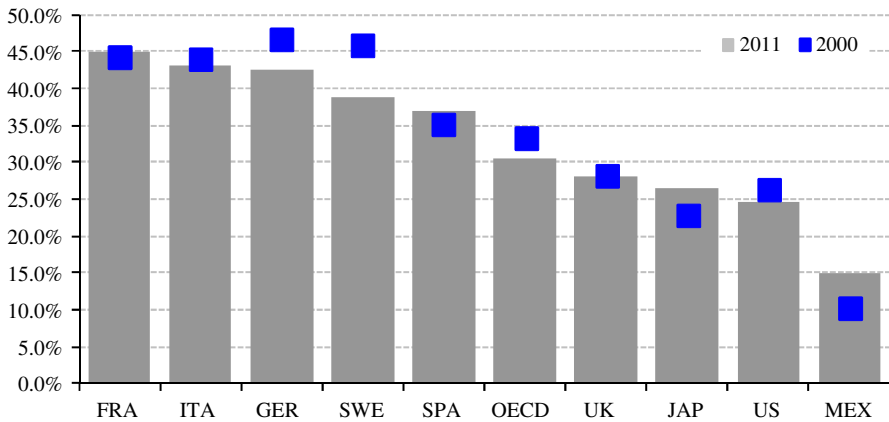


Fig. 1 Labour taxation in selected OECD countries (total average tax wedge, percentage of labour costs). Source: OECD Taxing Wages. Note: Two-earner married couple one at 100 % of average earnings and the other at 67 % of two children

institutions. The [European Central Bank \(2008\)](#) documented the disincentives to work (particularly for low-income workers) stemming from high marginal tax rates. In a similar vein, [Keane \(2011\)](#) reviewed the literature on labour supply by gender. Once again, there appears to be a considerable controversy over their response to changes in wages and taxes. This is especially the case for men, due to differences in the measurement of wages and human capital (for women most studies find a large labour supply elasticity).

This is the appropriate field for the meta-analysis approach. In contrast to these narrative surveys, meta-analysis allows revising the relevant empirical literature in a more formal and objective manner. As summarised by [Stanley and Jarrell \(1989\)](#),² meta-analysis starts with the compilation of an exhaustive sample of literature and the choice of the dependent variable (in our case, the degree of backward shifting of labour taxes, proxied by the estimated elasticity of net wages to taxation). A general set of ‘moderators’, i.e. variables that reflect the quantitative and qualitative features of the different studies that could be influencing their results (theoretical model and sample, among others) is then selected and tabulated using dummy variables. The meta-regression of the dependent variable on these moderators can be used to quantifying the ‘true dependent variable’, that is, the consensus result of the empirical literature on the effect of taxes on wages after controlling for methodological differences. And also, and probably more important, meta-analysis permits to show which aspects of the modelling, data and econometric techniques are important, or not, for the estimates.

A sensible starting point is [Fuchs et al. \(1998\)](#). Based on a survey of economics departments at 40 leading US universities, the authors conclude that employers bear

² This technique is being increasingly applied in labour economics. See, among others, [Jarrell and Stanley \(1990\)](#) for the analysis of unions and wage gap, [Card and Krueger \(1995\)](#) for minimum wages and employment, [Nijkamp and Poot \(2005\)](#) and [European Commission \(2005\)](#) for real wage elasticity, [Longhi et al. \(2006, 2007\)](#) for immigration, wages and employment, [Evers et al. \(2006\)](#) for taxes and working hours and [Card et al. \(2010\)](#) for active labour policies evaluation. By contrast, we are not aware of any meta-analysis applied to the economic incidence of either labour taxation or social contributions. See also [Stanley \(1998\)](#) for a meta-analysis of the Ricardian equivalence.

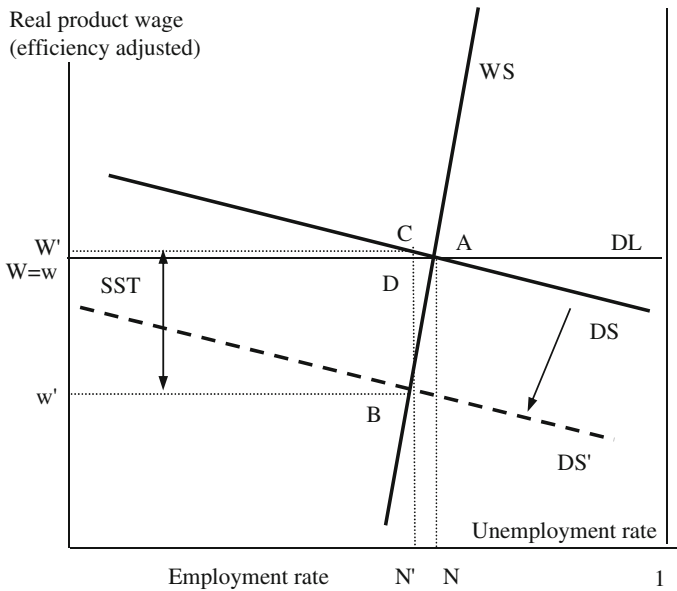


Fig. 2 Economic incidence of employer payroll taxes

20 % of firm social contributions, while employees bear the remaining 80 % via lower net wages. In other words, the conventional wisdom, in line with public economics textbooks, is that the relatively higher rigidity of labour supply with respect to labour demand determines that the market adjustment is mainly concentrated on wages, and not on employment.³

Figure 2 generalises this framework, taking into account wage bargaining and firm market behaviour, as suggested by Layard et al. (1991). In this ‘wage-setting schedule’, real wage accepted by workers (adjusted for trend growth in labour efficiency) varies with the unemployment rate: the higher the unemployment rate, the lower real wages will be. The position of the wage-setting schedule is influenced by a number of structural characteristics, such as the degree of trade union power, the generosity of unemployment benefits, the stringency of employment regulations or the efficiency of the matching process and, related to it, the centralisation of the wage-employment bargaining. The empirical literature confirmed the role of the centralisation of wage bargaining and union presence. Centralised economies with strong unions such as the Nordic countries, or decentralised wage bargaining with weak unions as the Anglo-Saxon exhibit better performance than the Continental and Mediterranean. We will

³ This degree of shifting is coherent with the labour supply and demand elasticities reported in the survey: 0.15 and -0.50 respectively. In partial equilibrium, shifting can be calculated as the ratio $0.50/[0.15 - (-0.50)]$, so that the implied estimate is 0.77 (see Fullerton and Metcalf (2002)). Therefore, employees would bear 77 % of social security tax burden. Roughly speaking, the simulations by Coenen et al. (2008) imply a shifting of nearly 86 %, since real wages increase 12.7 % following the reduction by 14.8 percentage points in firm social security contributions. A similar exercise calibrated for the Spanish economy by Boscá et al. (2009) reduces the share borne by firms to very close to the stylized figure (76 %: a 15 % cut in firm contributions leads to an increase of 11.4 % of real wages).

develop this issue afterwards. In this context, an increase in labour taxation (e.g. employer social security taxes, SST in the Fig. 2) would generate a downwards shift of the labour demand curve ('price-setting schedule'). Market equilibrium would move from A to B, generating a limited negative effect on the employment rate (from N to N'), since labour costs hardly increase (from W to W'). Workers would bear the major part of the tax burden (BD over BC), since net real wages fall from w to w' .

The main goal of this paper is to test whether this estimate of the economic incidence of labour taxation is consistent with the empirical literature on the subject. To do so, we quantify the effect of the different methodological approaches, and temporal and geographical coverage. We think we are the first to perform a meta-analysis of the incidence of labour taxes and social contributions. We place more weight on the methodological variables that stem from economic theory or from generally accepted empirical results, such as those related to nominal rigidities, the wage bargaining characteristics or the pension system design.

The paper is organized as follows. In Sect. 2, we present a brief description of the empirical literature of taxation and wages/labour costs, with a focus on social contributions. In Sect. 3, we report the basic methodology and the results of the meta-analysis regressions. They are grouped under three headings: 'basic moderators' that capture country and temporal fixed effects, 'economic moderators' to control the impact of the most relevant economic features and 'other moderators', mainly reflecting the econometric techniques. We also perform various robustness checks, based on the sample of estimates, the econometric methods and the procedure of specification selection. Finally, in Sect. 4 we present the main conclusions and general economic policy recommendations.

2 A brief survey of the empirical literature

2.1 Basic methodology

Public economics literature highlights two alternative definitions of the nature of social contributions. On the one hand, they can represent a deferred salary if the link between social contributions and social benefits (pensions and unemployment benefits) is high. This is the case of pension systems with defined contributions. On the other hand, they might be considered a regular payroll tax (such as the labour income tax), if the tax-benefit link is weak, as is the case in many defined-benefit pay-as-you-go pension systems. In this paper we opt for an intermediate position, so social contributions would be treated as a particular labour tax whose revenues are partially affected by social security financing. Therefore, we will concentrate on the labour market dynamics to analyse the incidence of social contributions and, in general, of labour taxes.

The focus will be placed on their economic incidence, i.e., who effectively bears the tax burden and what are its main economic effects (on wages, labour costs and equilibrium employment), rather than on the legal incidence.⁴ This involves analysing

⁴ Additionally, public economics literature defines differential incidence as that in which analysis is focused on the economic effects of substituting one tax by another, leaving total revenues constant. Finally, distributive incidence focuses on the income distribution effects of taxes.

the change in behaviour of economic agents after the establishment of the tax. As mentioned above, in the case of social contributions the literature points at two basic shifting processes. Consumers may bear the burden if firms have enough market power to adjust prices to the new tax rates. This is known as forward shifting. Alternatively, employees may bear the burden if the firm offsets the new tax burden with a (net) negative wage variation, leaving labour costs constant, a process known as backward shifting.

Therefore, the benchmark equation in the studies takes the form of a general wage equation

$$\mathbf{f}(\mathbf{W}_{it}) = \mathbf{g}(\mathbf{P}_{it}, \mathbf{Y}_{it}/\mathbf{N}_{it}, \mathbf{U}_{it}, \mathbf{h}(\mathbf{TAX}_{it}), \mathbf{X}_{it}) \quad (1)$$

where \mathbf{W} stands for the net nominal wage or the nominal labour cost (transformed in logs), \mathbf{P} for the price level (the output or the consumption deflator), \mathbf{Y}/\mathbf{N} is the labour productivity, \mathbf{U} the unemployment rate, \mathbf{TAX} stands for the tax wedge (again in logs), and \mathbf{X} is a set of control variables. Focusing on the latter, theoretical and empirical literature shows that tax incidence depends not only on competitive factors (such as labour supply and demand elasticities, and factor substitutability⁵), but also on the effect of economic institutions such as the degree of wage bargaining centralisation and union power, employment protection legislation, unemployment benefits, minimum wages and social security fairness (we will develop these issues in Sect. 3). Finally, functions \mathbf{f} and \mathbf{h} denote levels or growth rates (long or short run results), and \mathbf{g} is generally a linear function of the regressors. Every variable is specified for country or population group \mathbf{i} , and time \mathbf{t} .

Therefore, the dependent variable is the net wage elasticity to taxes. As previously stated, this figure proxies the degree of backward shifting. A -1.0 coefficient represents the full shifting scenario, where workers bear 100 % of taxes. A null elasticity implies that workers do not bear even part of the tax (null shifting). Intermediate results are more frequent and imply partial shifting processes.⁶

2.2 Database

We have assembled a database of 670 estimates of the impact of taxes on net wages from 52 papers, covering most of the OECD countries and some Latin American economies.⁷ This sample is based on a narrative survey (Melguizo 2009), updated

⁵ See Fullerton and Metcalf (2002) for a complete revision of tax incidence.

⁶ Sometimes, the dependent variable of equation 1 is defined as the labour cost, that is, inclusive of employer social security contributions, $W = w(1 + SST)$. In this case, tax elasticities would range between 0.0 (equivalent to -1.0 in the wage equation, if there is a full internal compensation of wages and taxes that leave labour costs constant), and 1.0 (equivalent to 0.0 in wage terms, when a tax increase fully impacts labour costs). We will use this equivalence in the next section to standardize our sample.

⁷ Latin American economies are particularly interesting due to the implementation of structural pension reforms (starting in 1981 in Chile, and followed in the mid-1990s in Argentina, Colombia, Mexico and Peru among others), which legally changed the nature of social security contributions from being a tax to a deferred salary.

using the IDEAS search engine.⁸ We normalised the results to the wage elasticity (based on the labour cost-wage equation equivalence explained in footnote 6), and directly chose this variable as the dependent one.

We restricted our core sample to the preferred estimate in each study for each country or region covered, based on the author's judgement (when absent, we use the usual statistic tests). This limited the sample to 124 observations from the 52 papers since some of the papers report estimates for different regions (labelled 'baseline sample'). This option remains open in the meta-analysis literature. Some authors claim that excluding some estimates may bias the studies, weakening the main advantage of meta-analysis over other surveying techniques. However, the inclusion of all the estimates is not cost-free: it may bias the analysis towards those studies that report more results for each country (if their results are similar), or may increase the variability of the results even if the authors manifest their preference for certain estimates (if their results vary).

Table 1 presents the list and definition of the moderators, while Table 2 summarises the main statistics of this baseline sample. On average, a 1.0 % increase in taxation reduces wages by 0.66 %. Therefore, overall employees bear two thirds of social security contributions. However, consistent with the mixed results of narrative surveys, studies display a very high dispersion of estimates, ranging from 0.91 to -2.54 . This significant dispersion can be due to differences in the statistical methods or the temporal and geographical coverage, or can stem from different institutional designs. Depending on the main methodological alternatives, the highest degree of shifting is obtained in cross-section analysis and in book publications and mimeos, where the elasticity is around -0.90 . By contrast, studies focused on Continental and Mediterranean economies or published in journals tend to limit the elasticity to -0.50 approximately.

As mentioned above, the empirical literature highlights the role of economic institutions, particularly the centralisation of wage bargaining and union presence. Centralised economies with strong unions or decentralised wage bargaining with weak unions exhibit lower unemployment rates. In addition, public sector effectiveness may compensate the disincentives from a high taxation, since workers perceive the tax-benefit linkage. Broadly speaking, Nordic and Anglo-Saxon economies have the systems with the best equilibrium, while Continental and Mediterranean economies tend to be placed in an unfavourable intermediate position.

Figure 3 supports this hypothesis suggesting that Nordic economies, characterised by their high centralisation, strong trade unions and effective governments are different. The mode estimate is full shifting (-1.0 elasticity), while in Anglo-Saxon, Continental and Mediterranean economies workers seem to bear half of the tax burden (-0.5).⁹ Are these differences statistically significant? Do the a priori results hold when controlling

⁸ <http://ideas.repec.org>. To be precise, in August 2009 we searched for references that contained either in the title, among the keywords, or in the abstract the terms: 'social contributions', 'social security taxes', 'labour taxes', 'social taxes', 'payroll taxes', 'incidence social contributions', 'incidence social security taxes', 'incidence labour taxes', 'incidence social taxes', 'who bears', 'tax burden', 'wages and taxes', 'labour costs and taxes', 'tax shifting', 'backward shifting' or 'forward shifting'.

⁹ See Figs. 4 and 5 in the Appendix for a parallel analysis depending on the tax wedge definition and short- and long-run results.

Table 1 Meta-regressors definition

| | |
|------------------------------|---|
| Economic incidence | = point estimate of net wage to taxation |
| Meta-regressors ^a | |
| Publication | |
| JOUR | = 1 ...if a study is published in a journal |
| BOOK | = 1 ...if a study is a book or a chapter in a book |
| WP | = 1 ...if a study is published as a working paper |
| MIM | = 1 ...if a study is Unpublished |
| Economic model ^b | |
| ANGLO | = 1 ...if a study covers Anglo-Saxon economies |
| CONT | = 1 ...if a study covers Continental or Mediterranean economies |
| NORD | = 1 ...if a study covers Nordic economies |
| Social security ^c | |
| BISSS | = 1 ...if a study refers to contributory systems ('Bismarckian') |
| BEVSS | = 1 ...if a study refers to redistributive systems ('Beveridge') |
| Fiscal wedge | |
| SALWEDGE | = 1 ...if a study defines taxation as salary wedge (prices and all taxes) |
| FISCWEDGE | = 1 ...if a study defines taxation as fiscal wedge |
| SALFISC | = 1 ...if a study defines taxation as salary or fiscal wedge |
| DIRTAX | = 1 ...if a study defines taxation as direct/labour fiscal wedge |
| CONTRIB | = 1 ...if a study defines taxation as social contributions |
| Focus | |
| DEPLEV | = 1 ...if results refers to the long run |
| DEPVAR | = 1 ...if results refers to the short run |
| Coverage | |
| ECO | = 1 ...if a study covers the entire economy |
| PRIVECO | = 1 ...if results are restricted to the market economy |
| Data | |
| TIME | = 1 ...if a study uses time-series data |
| Focus | |
| DEPLEV | = 1 ...if results refers to the long run |
| DEPVAR | = 1 ...if results refers to the short run |
| Coverage | |
| ECO | = 1 ...if a study covers the entire economy |
| PRIVECO | = 1 ...if results are restricted to the market economy |
| Data | |
| TIME | = 1 ...if a study uses time-series data |
| CROSS | = 1 ...if a study uses cross-section data |
| PANEL | = 1 ...if a study uses panel data |
| Estimation method | |
| OLS | = 1 ...if a study uses ordinary least squares |
| IV | = 1 ...if a study uses instrumental variables |
| OTHER | = 1 ...if a study uses other estimators |

Table 1 continued

| | |
|---------------------|---|
| Economic incidence | = point estimate of net wage to taxation |
| Data frequency | |
| QUINQ | = 1 ...if a study uses quinquennial data |
| BIANNUAL | = 1 ...if a study uses biannual data |
| ANNUAL | = 1 ...if a study uses annual data |
| HALF | = 1 ...if a study uses half-yearly data |
| QUART | = 1 ...if a study uses quarterly data |
| Variable definition | |
| W | = 1 ...if a study defines the dependent variable as the net wage |
| LC | = 1 ...if a study defines the dependent variable as the labour cost |

^a Additionally, we controlled for geography (OECD, US, Latin America and Spain), for decades, and for other labour market institutions included in the specifications (unions, minimum wage, unemployment benefits)

^b Based on [Alesina and Perotti \(1997\)](#), [Daveri and Tabellini \(2000\)](#) and [Sapir \(2006\)](#), and own elaboration. ANGLO includes Australia, Canada, Chile, Ireland, Japan, New Zealand, Switzerland, UK and US. CONT includes Austria, Belgium, Colombia, France, Germany, Greece, Luxembourg, Italy, Mexico, Netherlands, Portugal and Spain. NORD includes Denmark, Finland, Norway and Sweden

^c Based on [Disney \(2004\)](#), and own elaboration. BEVSS includes Australia, Canada, Denmark, Ireland, Japan, Netherlands, New Zealand, Switzerland, UK and US. BISSS includes Austria, Belgium, Chile, Colombia, Finland, France, Germany, Greece, Luxembourg, Italy, Mexico, Norway, Portugal, Sweden and Spain

for the complete set of methodological differences? The aim of the meta-analysis performed in the next section is precisely to answer these questions quantitatively.

3 Meta-regression analysis

3.1 Meta-analysis approach

Given the relatively recent arrival of the meta-analysis technique in economics, there is no standardised empirical strategy. As previously explained, we restricted the sample to 124 observations (from a complete sample of 670 observations), although we will use the complete dataset to test the robustness of the results. Due to the presence of some extreme values in the sample, robust regression estimation is recommended.¹⁰ We will also include some weighted least squares (WLS) estimates to control the quality of the studies.¹¹

¹⁰ As programmed in Stata 11 using ‘rreg’. We also estimated the specifications using quantile regressions (‘qreg’), obtaining very similar results. Estimations are available upon request.

¹¹ In a previous version of the paper, we opted for ordinary least squares (in line with [Stanley and Jarrel \(1989\)](#)), controlling the effects of four statistically identified outliers ([Brittain 1972](#); [Argimón and González-Páramo 1987](#); [Anderson and Meyer 1998](#); [Hamaaki and Iwamoto 2008](#)). As an additional robustness check, we estimated most of the specifications truncating the elasticity values to those set in the economic theory, i.e. from 0.0 to 1.0. In both cases, results are very similar to those reported in the main text. Estimations are available upon request.

Table 2 Taxation and wages (selected descriptive statistics)

| | Moderator | No. | Mean | SD | Max | Min |
|-----------------|-----------------|------------|--------------|-------------|-------------|--------------|
| Total | | 124 | −0.66 | 0.51 | 0.91 | −2.54 |
| Publication | Journal | 58 | −0.52 | 0.49 | 0.41 | −1.60 |
| | Book/Chapter | 21 | −0.89 | 0.52 | 0.06 | −2.54 |
| | Working paper | 41 | −0.71 | 0.51 | 0.91 | −2.12 |
| | Mimeo | 4 | −0.90 | 0.13 | −0.73 | −1.00 |
| Economic model | Anglo-Saxon | 57 | −0.65 | 0.60 | 0.40 | −2.54 |
| | Continental-Med | 41 | −0.54 | 0.45 | 0.91 | −1.26 |
| | Nordic | 27 | −0.75 | 0.36 | 0.17 | −1.00 |
| Social Security | Bismarckian | 59 | −0.62 | 0.43 | 0.91 | −1.26 |
| | Beveridge | 65 | −0.65 | 0.58 | 0.40 | −2.54 |
| Fiscal wedge | Salary wedge | 27 | −0.79 | 0.30 | 0.06 | −1.26 |
| | Fiscal wedge | 6 | −0.69 | 0.38 | 0.00 | −1.00 |
| | Direct taxes | 22 | −0.64 | 0.42 | 0.41 | −1.00 |
| | Social taxes | 69 | −0.61 | 0.60 | 0.91 | −2.54 |
| Focus | Short-run | 38 | −0.60 | 0.49 | 0.40 | −2.07 |
| | Long-run | 86 | −0.68 | 0.52 | 0.91 | −2.54 |
| Coverage | Total economy | 64 | −0.69 | 0.39 | 0.91 | −1.26 |
| | Private economy | 60 | −0.63 | 0.62 | 0.41 | −2.54 |
| Approach | Time series | 84 | −0.64 | 0.48 | 0.91 | −2.54 |
| | Cross-section | 3 | −0.92 | 0.49 | −0.41 | −1.39 |
| | Panel | 37 | −0.68 | 0.59 | 0.41 | −2.12 |
| Geography | LatAm | 7 | −0.50 | 0.40 | −0.14 | −1.12 |
| | Spain | 12 | −0.43 | 0.60 | 0.91 | −1.00 |
| | US | 32 | −0.68 | 0.65 | 0.40 | −2.54 |
| | OECD | 119 | −0.67 | 0.52 | 0.91 | −2.54 |

We opt for a sequential approach, in line with [Knell and Stix \(2003\)](#), [Evers et al. \(2006\)](#) and [Card et al. \(2010\)](#). Firstly, we test the significance of geographical and temporal (by decades, from the 1950s to the 2000s) moderators, in the spirit of panel-data fixed effects estimation. This enables a basic specification to be obtained, in which we will include the main economic and methodological moderators as a second step (see [Table 1](#), for the complete set). Instead of including all of these moderators simultaneously, we opt to include them one by one to avoid multicollinearity problems. Afterwards, we incorporate the significant moderators in a combined specification. As a final check, we perform an extreme-bounds analysis to evaluate whether results are influenced by the sequence of estimations.

In what follows, we will concentrate on reporting the significant results and focus on the role of economic moderators, which naturally suggest economic policy

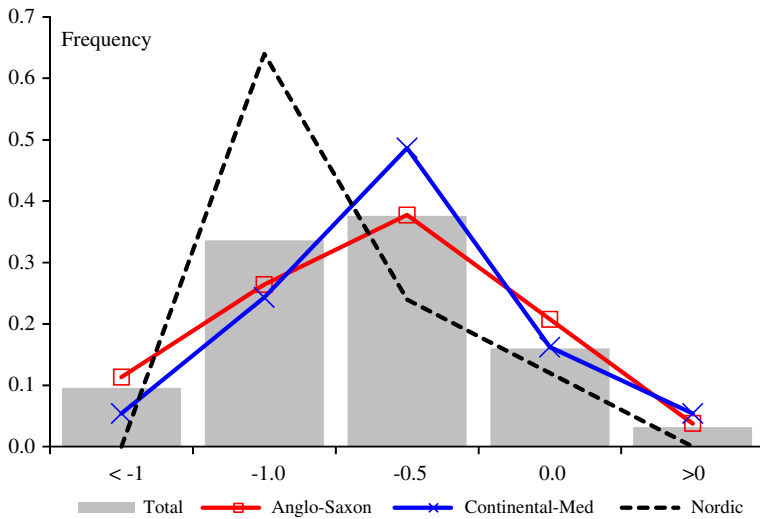


Fig. 3 Distribution of wage elasticity to taxes (elasticity of net wages to taxation)

recommendations.¹² The meta-regression specification, based on Eq. 1, is straightforward

$$\beta_s = \mathbf{b} + \alpha_{s,i} + \alpha_{s,t} + \gamma \mathbf{X}_s + \mathbf{u}_s \quad (2)$$

where β_s is the reported wage elasticity to taxes in study s (defined as the impact of taxes on net wages), \mathbf{b} is the ‘true’ elasticity, $\alpha_{s,i}$ and $\alpha_{s,t}$ are vectors with geographical and temporal ‘fixed effects’, \mathbf{X}_s is a vector of moderator variables, and \mathbf{u}_s is the error term. The independent variables are represented by dummies (for instance, if a study sample covers the period 1980–2008, D50 to D70 will be marked as a column of 0s for the particular study, while D80 to D00 will be represented by a column of 1s). \mathbf{X}_s will basically reflect imperfect competition variables stemming from the theoretical and empirical literature, such as the public sector effectiveness, labour market institutions (unions, wage bargaining), social security rules and tax wedge mix.

3.2 Results

Baseline specification (‘fixed effects’ moderators) Firstly, we aim to obtain a baseline specification in the spirit of the fixed-effect panel estimation. Alternative estimations rejected robust geographical effects (we tested for differences in studies covering Latin America, Spain and the US vs. other OECD economies). For this reason, and since some of the economic moderators may be correlated, the basic specification does not include them. No significant temporal fixed effects, defined by the sample period in the data used (identified by the decades covered in case of time series analysis), are found either.

¹² See Table 6 for additional specifications. The complete set of estimates is available upon request.

Another issue that should be controlled from the beginning in order to avoid spurious results is the presence of publication bias. It may arise from the tendency to report and/or to publish only the significant results, rejecting the null hypothesis of no effect. In order to reduce its potential influence, we included in our meta-database both published and unpublished papers. More than one third of the sample, 45 estimates, comes either from working papers or mimeos (see Table 2).¹³

Based on the previous elements, in this initial benchmark specification the wage elasticity to taxes is -0.65 , so workers bear almost two thirds of the tax burden (column I in Table 3).

Economic moderators The attention devoted to the economic effects of labour taxation from international organisations, policy makers and academia ranks among the highest in applied macroeconomics, generating an enormous amount of research. However, the main issues affecting our results can be restricted to the following: the role of taxes under different institutional settings, the tax wedge composition, short- vs. long-run results and the role of the social security scheme.

Economic institutions matter. The theoretical and empirical literature shows that the impact of labour taxes on labour costs and unemployment is higher in economies with an intermediate centralisation of the wage bargaining process and a strong trade union presence (see Calmfors and Driffill 1988; Alesina and Perotti 1997; Daveri and Tabellini 2000). In this context, common to Continental and Mediterranean countries, the discipline effect of competition or the internalization of externalities are weak. This contrasts with the behaviour in decentralized Anglo-Saxon economies, and unionized centralized Nordic countries. Our analysis partially confirms this hypothesis. Column II in Table 3 shows that the average elasticity is -0.59 . However, the degree of shifting is much higher in Nordic countries at -0.79 ($-0.59 - 0.20$). This could be an indicator of the benefits of good governance, since Nordic public sectors are among the most effective (see Sapir 2006). By contrast, no significant differences were found between Anglo-Saxon and Continental-Mediterranean economies.

Throughout this paper we use the terms ‘tax wedge’, ‘labour taxation’ and ‘social security contributions’ in a flexible way, implicitly accepting the tax invariance theorem. However, there are several reasons to justify a differential impact of the fiscal wedge components. Focusing on social contributions, even though they are usually a payroll tax, their tax base and tariff usually differ from those of personal income taxes (not to mention the linkage effect), and consumption taxes (see OECD 1990, 2007). Additionally, the salary wedge includes the price wedge, that is, the gap between producer and consumer prices.

Formally, we define pit as the personal income tax effective rate, $essc$ the employee social security contributions rate, $fssc$ the firm social security contributions rate, ct the

¹³ As an additional way to control for the possible publication bias, we included as an additional regressor the standard error of the estimates. Unfortunately, its reporting is neither unanimous (50 out of the 124 results in our meta-database did not include them), neither homogeneous (some authors report standard errors, some others robust standard errors). In any case, we carried out the estimations, and found out mixed results. As shown in Table 6, robust regression estimation (specification VII) results suggest its presence, while WLS (specification XVII) rejects it. For a detailed explanation of this issue and other methodologies to test it based on the sample size, see Ashenfelter et al. (2000).

Table 3 Main results

| Equations | Moderators ^a | I | II | III | IV | V | VI |
|-----------------|-------------------------|-----------------|-----------------|-----------------|-----------------|-----------------|-----------------|
| Incidence | C | −0.65*** (0.04) | −0.59*** (0.06) | −0.59*** (0.05) | −0.74*** (0.05) | −0.52*** (0.07) | −0.70*** (0.05) |
| Economic model | NORD | | −0.20* (0.11) | | | −0.12 (0.09) | −0.18* (0.10) |
| | CONT | | −0.04 (0.10) | | | | |
| Fiscal wedge | SALFISC | | | −0.20** (0.10) | | −0.07 (0.11) | |
| Short run | DEPVAR | | | | 0.31*** (0.09) | 0.24*** (0.08) | 0.31*** (0.09) |
| Coverage | ECO | | | | | −0.02 (0.09) | |
| Dependent | LC | | | | | −0.29** (0.10) | |
| <i>N</i> | | 124 | 124 | 124 | 124 | 124 | 124 |
| <i>F</i> | | 0.00 | 1.48 | 4.31 | 6.39 | 6.39 | 7.92 |
| Prob > <i>F</i> | | – | 0.23 | 0.04 | 0.00 | 0.00 | 0.00 |

Standard errors in parentheses ***, **, * denote significance at 1, 5 and 10 % levels ^a Default specification: Long-run labour taxation effects in an OECD/Latin American market economy, with a Bismarckian pension system

consumption tax rate, pp the domestic output deflator, pp^* international prices deflator, a the share of domestic products in the consumption basket, and e the exchange rate. The ‘salary wedge’ (wedge between the nominal labour costs and the net real wage) can be defined as $\left[\frac{(1+ct)(1+fssc)}{(1-pit)(1-essc)} \right] \left[\frac{epp^*}{pp} \right]^{1-a}$. The ‘fiscal wedge’ includes exclusively the tax components of the salary wedge and can be expressed as $\left[\frac{(1+ct)(1+fssc)}{(1-pit)(1-essc)} \right]$. The ‘direct tax wedge’ only focuses on taxes on income as $\left[\frac{(1+fssc)}{(1-pit)(1-essc)} \right]$ and, finally, the ‘employer social taxes wedge’ can be written as $(1+fssc)$.

Therefore, we tested the impact of alternative fiscal wedge definitions, starting from the general salary wedge (fiscal wedge plus price wedge), down to employer social security contributions. Results confirm that different taxes have different effects on wages. While the elasticity of net wages to labour taxes is -0.59 , the salary and fiscal wedge impact (inclusive of indirect taxes and of price wedge between output and consumption in the first case) is close to 80 % ($-0.79 = -0.59 - 0.20$, as shown in column III in Table 3).

Moreover, as highlighted in Hamermesh (1993), the presence of lags in the shifting of social contribution is usually significant (on average, long-run shifting equilibrium takes more than a year) due to the presence of nominal rigidities. Our estimation confirms this point (column IV in Table 3). The short-run elasticity $-0.43(-0.74 + 0.31)$ is lower than the long-run counterpart, -0.74 . Therefore, workers bear less than half of the tax burden in the short run (generally represented by results during the year of the tax change).

Finally, a key institutional issue that has to be tested is related to the social security scheme. In particular, the literature highlights a potential role of the ‘linkage effect’, that is, the perception of the link between contributions and social benefits (see Gruber and Krueger 1990; Gruber 1994a,b; Disney 2004). Social contributions would have no effect on the equilibrium employment rate if agents perceive a full linkage effect, since they become a deferred salary (positively shifting the labour supply curve). This tends to be the case in contributory *Bismarckian* social security systems, in contrast to redistributive pension systems *à la Beveridge*. Our empirical analysis obtains the expected sign, since shifting is lower in more redistributive systems, but the effect is not significant (coefficient 0.10, column VIII in Table 6).¹⁴

Other moderators A second set of moderators tries to reflect methodological differences, from both the data (economy-wide or private sector, and its frequency) and the estimation techniques (cross-section, time series or panel, and selected estimator). Our analysis suggests that data coverage is relevant. Studies that refer to the entire economy, and not just the market economy, tend to show higher negative elasticities, $-0.73(-0.52 - 0.21)$; column IX in Table 6). We are not convinced of its implications, since wage bargaining in the public sector is not competitive. Finally, more than half

¹⁴ The absence of gains from an increase in the actuarial fairness of the social security (i.e. lower redistribution) may stem from a divergence between the institutional design and its perception. Boeri et al. (2001) highlight a significant lack of knowledge on basic social security institutional issues in the main economies of the EU (e.g. on the basic pay-as-you-go functioning, or even on the tax rates). In this context, a theoretically more efficient institutional setting may not fully generate its potential benefits.

of the papers include labour costs, rather than the net wage, as the dependent variable. Those studies tend to obtain a higher negative elasticity ($-0.80 = -0.43 - 0.37$; column X in Table 6), although no economic implication seems straightforward either.

Combined results To conclude, we incorporated all the individually significant moderators into the specification (column V in Table 3). Every coefficient keeps its sign, but only the temporal focus and the dependent variable definition remain significant. Finally, column VI (Table 3) presents our baseline specification, focused on the moderators that have a more solid economic foundation. In this case, the net wage elasticity is -0.70 , so workers bear 70 % of the tax burden via lower wages (or lower wage increases), somewhat lower than the starting estimate surveyed in Fuchs et al. (1998). In the Nordic economies the degree of shifting is higher, close to 90 % ($-0.88 = -0.70 - 0.18$), so nearly all the tax changes are offset by negative wage variations. Finally, we confirm that shifting takes time, and in the short run workers bear less than 40 % of the tax burden ($-0.39 = -0.70 + 0.31$).¹⁵

3.3 Robustness checks

In order to further check the robustness of the results, we first tested whether the main results hold for the whole sample (670 elasticities from the 52 papers). Results, reported in columns XI to XV in Table 6, are in line with the ones reported for the core sample. These additional tests confirm that shifting is higher in Nordic economies, in studies that include indirect taxes in the tax wedge and in the long run. However, only the latter is significant. The publication bias hypothesis seems also present, with the aforementioned caveats on its downsizing effects to the sample. Taking all these considerations on board, our baseline specification for the complete sample is reported in column XIV. The net wage elasticity is -0.85 , so workers bear 85 % of the tax burden. In the short term the backward shifting process is lower as well ($-0.51 = -0.85 + 0.34$).

Additionally, we performed a quality control study of the literature other than the publication format (books, journals, working papers and mimeos). Specifically, we used *Google Scholar* to weight each estimate for the number of citations.¹⁶ The weighted least square estimation of the baseline sample is reported in columns XVI to XVIII in Table 6. In comparison to our baseline specification, estimates confirm that taxation shifting is significantly lower in the short run. However, the shifting degree seems higher, and does not depend on the economic model.

¹⁵ The selected specification passed the basic econometric tests of heteroskedasticity and autocorrelation, once we ordered the papers by year of publication at the conventional significance levels.

¹⁶ <http://scholar.google.com>, accessed December 24, 2009. The 52 papers accumulated 5.927 citations, Alesina and Perotti (1997), Gruber (1994b, 1997) and Layard et al. (1991) being the most popular. The weight factor is the share of its citations on the total amount, divided by the number of reported preferred estimates. Alternatively, other authors use the citations reported in the *ISI Web of Knowledge* (see, for instance, European Commission 2005). However, this source restricts the sample to papers published in JCR journals to almost half of them, excluding some of the most popular references. In any case, the correlation between the alternative weighting factors is quite high. See Table 5 in the Appendix for more details.

Finally, we tested how sensitive results are depending on the specification selection procedure. In particular, we calculated the confidence intervals of the impact of taxes on wages, based on the ‘extreme-bounds analysis’.¹⁷ The lowest bound is obtained including fixed effects and controlling for the study coverage (ECO) and the publication format (BOOK, WP + MIM): -0.10 (-0.45 minus two standard errors). The highest one stems from a specification with the social security model (BEVSS) and the time horizon (DEPVAR): -1.09 (-0.77 plus two standard errors). Therefore, this test robustly rejects the no-shifting hypothesis.

4 Conclusions

In this paper we have applied the meta-regression technique to analyse the results from the empirical literature on the economic incidence of labour taxes and social contributions. In particular, we have focused on the effects of taxation on wages to find out whether employees bear the tax burden due to lower net wages. This is a relevant question, due both to the significant dispersion of the results found in the literature and to its economic policy implications.

We have based our empirical analysis on an original database of 52 empirical papers, from which we extracted 124 estimates. On average, in the core sample a 1.0 % increase in taxation reduces wages by 0.66 %. Therefore, in line with the literature on distributive incidence, employees would bear nearly two thirds of social security contributions. However, the literature exhibits a very high degree of variability.

The meta-regression analysis suggests that this figure is affected by basic economic institutions (summarized in three models, namely Anglo-Saxon, Continental-Mediterranean and Nordic), and by the tax wedge definition (in particular, the inclusion of indirect taxes). In our preferred specification, the elasticity of wages to taxes is -0.70 in the default option (i.e. a non-Nordic economy in the long run). Therefore, workers bear 70 % of taxes. In the Nordic economies the degree of shifting is close to full (-0.88), so all tax changes are almost entirely offset by a wage variation. Moreover, the impact of taxes on wages differs in the short term. The degree of shifting is much lower in the short run: workers bear less than half of the tax burden. Finally, although not included in our preferred specification, consumption taxes may be more prone to shifting.

The robustness checks carried out using the complete set of estimates, alternative estimators and extreme-bounds analysis confirm this result. However, despite the numerous set of controlling variables, a significant part of the differing conclusions found in the empirical literature studied remains unexplained. A first extension would

¹⁷ Incidence and empirical growth analysis are both exposed to multicollinearity problems, given the variety of relevant factors highlighted in the literature. This makes difficult to choose its key determinants from exhaustive specifications. Results may also be biased in reduced specifications depending on the inclusion order of the regressors. [Levine and Renelt \(1992\)](#) proposed an automatic procedure to analyse the determinants of growth, based on computing the ‘extreme bounds’ computed from all the combinations of regressors. We concentrated on the economic incidence parameter (coefficient b in Eq. 2), estimating 21 specifications, which are available upon request.

be to estimate the combined effect of the moderators. Additionally, future research may focus on the heterogeneity (by age, sex and skill of the workers, or by industries) or on the direct effect of taxes on unemployment.

With conventional caution, some potential economic policy implications can be drawn from these results. Firstly, policy makers should take into account that even if higher social security contributions have a limited effect on employment in the long run, these are not entirely negligible, especially in the short run. This fact poses limits to strengthening social protection systems via higher revenues. Secondly, intermediate results pointed to significant employment gains from a revenue-neutral tax reform, increasing consumption taxes and lowering labour taxes. Obviously, this tax shift might also involve some problematic aspects, namely a short-term inflationary effect and a change in income distribution, which should be evaluated. Finally, our estimates support that taxation in Nordic economies, characterized by a high coordination of wage bargaining and effective public sectors, is more employment-friendly.

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Appendix

See Tables 4, 5, 6 and Figs. 4, 5, 6.

Table 4 Meta-analysis database

| Paper | Publication |
|--|-------------------------------------|
| Alesina and Perotti (1997) | American Economic Review |
| Anderson and Meyer (1997) | Journal of Public Economics |
| Anderson and Meyer (1998) | NBER Working Paper |
| Argimón and González-Páramo (1987) | FIES Documentos de Trabajo |
| Arpaia and Carone (2004) | European Commission Economic Papers |
| Baicker and Chandra (2006) | Journal of Labor Economics |
| Bell et al. (2002) | Bank of England Quarterly Bulletin |
| Brittain (1971) | American Economic Review |
| Brittain (1972) | The Brookings Institution |
| Brunello et al. (2002) | CESifo Working Paper |
| Calmfors and Nymoen (1990) | Economic Policy |
| Cazorla and Madero (2007) | CONSAR Documento de Trabajo |
| Coe and Krueger (1990) | IMF Working Paper |
| Cox-Edwards (2002) | Stanford University Working Paper |
| Daveri and Tabellini (2000) | Economic Policy |
| Dolado et al. (1986) | Economica |

Table 4 continued

| Paper | Publication |
|---------------------------------|--|
| Dye (1985) | Eastern Economic Journal |
| Escobedo (1991) | Investigaciones Económicas |
| Estrada et al. (2002) | Moneda y Crédito |
| Forslund (1995) | Swedish Economic Policy Review |
| Franz and Gordon (1993) | European Economic Review |
| Gordon (1971) | Brookings Papers on Economic Activity |
| Griffith et al. (2006) | CEPR Discussion Paper |
| Griffith et al. (2007) | Economic Journal |
| Gruber (1994a) | American Economic Review |
| Gruber (1997) | Journal of Labor Economics |
| Gruber and Krueger (1990) | NBER Working Paper |
| Hamaaki and Iwamoto (2008) | University of Tokyo Working Paper |
| Hamermesh (1979) | Southern Economic Journal |
| Holmlund (1983) | Scandinavian Journal of Economics |
| Honkapohja and Koskela (1999) | Economic Policy |
| Hugues (1985) | The Economic and Social Research Institute |
| Kananassou et al. (2007) | Mimeo |
| Komamura and Yamada (2004) | NBER Working Paper |
| Kugler and Kugler (2003) | CEPR Discussion Paper |
| Layard et al. (1991) | Oxford University Press |
| Leuthold (1975) | Public Finance Quarterly |
| Murphy (2007) | Labour Economics |
| Muysken et al. (1999) | Applied Economics |
| Nunziata (2001) | Mimeo |
| OECD (1990) | OECD Employment Outlook |
| Ooghe et al. (2003) | Empirica |
| Pehkonen (1999) | Finnish Economic Papers |
| Perry (1970) | Brookings Papers on Economic Activity |
| Pissarides (1991) | Economica |
| Tachibanaki and Yokoyama (2008) | Japanese Economic Review |
| Tyrväinen (1995) | OECD Jobs Study Working Paper Series |
| Vaillancourt and Marceau (1990) | Economics Letters |
| Van der Horst (2003) | ENEPRI Working Paper |
| Vroman (1974a) | Public Finance |
| Vroman (1974b) | Applied Economics |
| Weitenberg (1969) | Public Finance |

Table 5 Top cited papers and weights

| Paper | Google Scholar | | ISI Web of Knowledge | |
|-------------------------------|----------------|------------|----------------------|------------|
| | Citations | Weight (%) | Citations | Weight (%) |
| Gruber (1994a) | 535 | 9.1 | 169 | 36.7 |
| Gruber (1997) | 225 | 3.8 | 12 | 2.6 |
| Layard et al. (1991) | 3,027 | 2.9 | n.a. | n.a. |
| Gordon (1971) | 139 | 2.4 | 31 | 6.7 |
| Gruber and Krueger (1990) | 130 | 2.2 | n.a. | n.a. |
| Honkapohja and Koskela (1999) | 113 | 1.9 | 1 | 0.2 |
| Brittain (1971) | 98 | 1.7 | 33 | 3.6 |
| Alesina and Perotti (1997) | 281 | 1.6 | 40 | 2.9 |
| Franz and Gordon (1993) | 92 | 0.8 | 23 | 2.5 |
| Perry (1970) | 45 | 0.8 | 36 | 7.8 |
| Daveri and Tabellini (2000) | 49 | 0.3 | 22 | 1.6 |
| Pissarides (1991) | 56 | 0.1 | 12 | 2.6 |

Note Weights are calculated as the percentage on citations, divided by the number of reported estimates

Table 6 Additional meta-analysis results

| Equations | Robust regressions: Baseline sample: Alternative moderators | | | | Robust regressions: Total sample | | | | WLS: Baseline sample | | | | |
|------------------|---|-----------------|-----------------|-----------------|----------------------------------|-----------------|-----------------|-----------------|----------------------|-----------------|-----------------|-----------------|-----------------|
| | Moderators ^a | VII | VIII | IX | X | XI | XII | XIII | XIV | XV | XVI | XVII | XVIII |
| Incidence | C | −0.50*** (0.07) | −0.70*** (0.06) | −0.52*** (0.06) | −0.43*** (0.05) | −0.54*** (0.03) | −0.70*** (0.02) | −0.71*** (0.02) | −0.85*** (0.03) | −0.83*** (0.03) | −1.22*** (0.17) | −0.90*** (0.12) | −0.68*** (0.10) |
| Publication bias | SD | −0.42** (0.16) | | | | −0.47*** (0.04) | | | | | | | 0.07 (0.43) |
| Economic model | NORD | | | | | −0.12* (0.09) | | | −0.09 (0.07) | −0.09 (0.13) | 0.01 (0.14) | | |
| Social security | BEVSS | | 0.10 (0.08) | | | | | | | | | | |
| Fiscal wedge | SALFISC | | | | | | | −0.11 (0.07) | | | −0.06 (0.23) | | |
| Short run | DEPVAR | | | | | | | | 0.34*** (0.04) | 0.34*** (0.04) | 0.45** (0.20) | 0.38* (0.21) | |
| Coverage | ECO | | | −0.21** (0.08) | | | | | | | 0.06 (0.22) | | |
| Dependent | LC | | | | −0.37*** (0.07) | | | | | | 0.46 (0.22) | | |
| <i>N</i> | 74 | 124 | 124 | 124 | 124 | 508 | 670 | 670 | 670 | 670 | 123 | 123 | 74 |
| <i>F</i> | 7.05 | 1.18 | 6.55 | 25.71 | 111.48 | 3.04 | 3.04 | 2.19 | 78.34 | 40.66 | 2.13 | 1.63 | 0.02 |
| Prob > <i>F</i> | 0.01 | 0.28 | 0.01 | 0.00 | 0.00 | 0.00 | 0.08 | 0.14 | 0.00 | 0.00 | 0.07 | 0.20 | 0.88 |

Standard errors in parentheses ***, **, * denote significance at 1, 5 and 10 % levels ^a Default specification: long-run labour taxation effects in an OECD/Latin American market economy, with a Bismarckian pension system

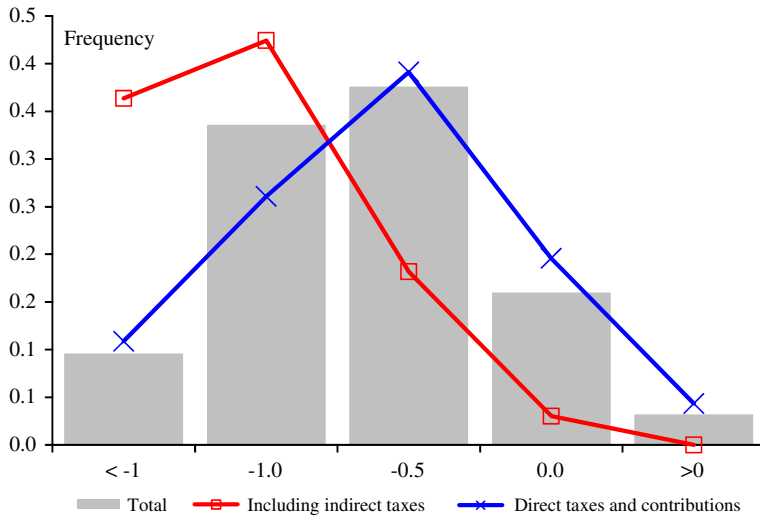


Fig. 4 Distribution of wage elasticity to taxes (tax wedge composition)

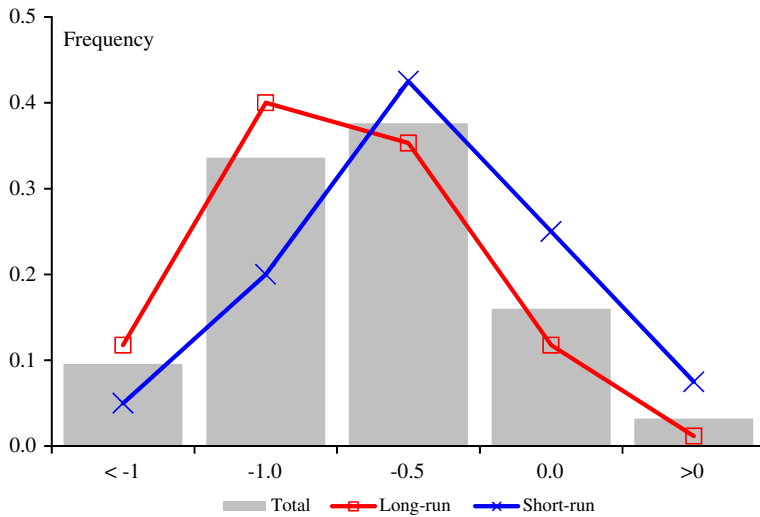


Fig. 5 Distribution of wage elasticity to taxes (long-run vs short-run results)

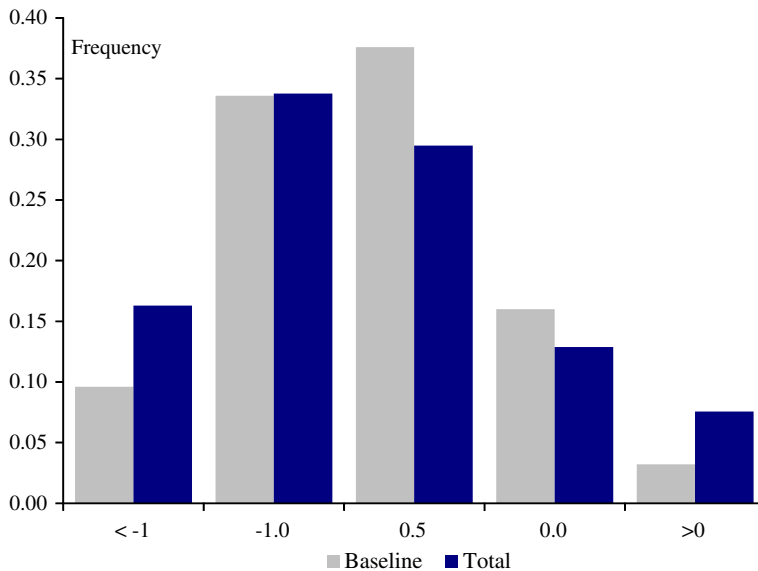


Fig. 6 Distribution of wage elasticity to taxes (total vs. baseline sample)

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